

mid-stochastic-fall2022-solution

November 2022

1 Q1

1.1 1

Obviously not!

1.2 2

ب) درست

Suppose that for a Poisson process at rate λ , we condition on the event $\{N(t) = 1\}$, the event that exactly one arrival occurred during $(0, t]$. We might conjecture that under such conditioning, t_1 should be uniformly distributed over $(0, t)$. To see that this is in fact so, choose $s \in (0, t)$. Then

$$\begin{aligned} P(t_1 \leq s | N(t) = 1) &= \frac{P(t_1 \leq s, N(t) = 1)}{P(N(t) = 1)} \\ &= \frac{P(N(s) = 1, N(t) - N(s) = 0)}{P(N(t) = 1)} \\ &= \frac{e^{-\lambda s} \lambda s e^{-\lambda(t-s)}}{e^{-\lambda t} \lambda t} \\ &= \frac{s}{t}. \end{aligned}$$

3 Q3

برش سوم

$$P(T_1, \dots, T_{n-1} | T_n)$$

$$= \frac{P(T_1, \dots, T_{n-1}, T_n)}{P(T_n)}$$

$$P(T_1, \dots, T_n) = \frac{P(T_n | T_1, \dots, T_{n-1}) P(T_1, \dots, T_{n-1})}{P(T_n)}$$

$$= e^{-\lambda(T_n - T_{n-1})} P(T_1, \dots, T_{n-1})$$

$$P(T_1, \dots, T_n) = e^{-\lambda T_n}$$

با استواری n به دست می آید:

$$T_n : \text{مجموع n متغیر تصادفی مستقل} \rightarrow T_n \sim \lambda^n T_n^{n-1} e^{-\lambda T_n}$$

$$\rightarrow P(T_1, \dots, T_{n-1} | T_n) = \frac{e^{-\lambda T_n}}{\lambda^n T_n^{n-1} e^{-\lambda T_n}} = \lambda^{-n} T_n^{-(n-1)}$$

$$= \lambda^{-n} T_n^{-(n-1)}$$

$$E[(T_n - T_{n-1}) + (T_n - T_{n-2}) + \dots + (T_n - T_1)]$$

$$= \sum_{i=1}^{n-1} E[T_n - T_i] = \sum_{i=1}^{n-1} \lambda(n-i) = \sum_{i=1}^{n-1} \lambda i$$

$$= \lambda \frac{n(n-1)}{2} = \lambda \frac{10 \times 9}{2}$$

$$\rightarrow \text{مقدار مورد نیاز} = \lambda \frac{n(n-1)}{2} = 45 \lambda$$

(13)

$$\begin{aligned}
 P(T_K \leq s | N_t = n) &= P(N_s \geq K | N_t = n) \\
 &= \sum_{i=K}^n P(N_s = i | N_t = n) = \sum_{i=K}^n \frac{P(N_s = i, N_{t-s} = n-i)}{P(N_t = n)} \\
 &= \sum_{i=K}^n \frac{e^{-\lambda s} \frac{(\lambda s)^i}{i!} \times e^{-\lambda(t-s)} \frac{(\lambda(t-s))^{n-i}}{(n-i)!}}{\lambda t^n e^{-\lambda t}} \\
 &= \sum_{i=K}^n \frac{s^i (t-s)^{n-i}}{t^n} \binom{n}{i} = \sum_{i=K}^n \left(\frac{s}{t}\right)^i \left(1 - \frac{s}{t}\right)^{n-i} \binom{n}{i} \\
 \rightarrow P(T_K \leq s | N_t = n) &= \sum_{i=K}^n \left(\frac{s}{t}\right)^i \left(1 - \frac{s}{t}\right)^{n-i} \binom{n}{i}
 \end{aligned}$$

(14)

$$\begin{aligned}
 E\left[\sum_{k=1}^n (t - T_k) | N_t = n\right] &= \sum_{k=1}^n E[t - T_k | N_t = n] = \sum_{k=1}^n \int_0^t P(T_k \leq s | N_t = n) ds \\
 &= \sum_{k=1}^n \int_0^t \sum_{i=k}^n \left(\frac{s}{t}\right)^i \left(1 - \frac{s}{t}\right)^{n-i} \binom{n}{i} ds \\
 &= \int_0^t \sum_{k=1}^n k \left(\frac{s}{t}\right)^k \left(1 - \frac{s}{t}\right)^{n-k} \binom{n}{k} ds = \int_0^t \sum_{k=1}^n k p^k (1-p)^{n-k} \binom{n}{k} dp \\
 &= \sum_{k=1}^n k \binom{n}{k} \left(\frac{(n+1)!}{k!(n-k)!}\right)^{-1} = \frac{t}{n+1} \sum_{k=1}^n k = \frac{tn}{2}
 \end{aligned}$$

Scanned with CamScanner

4 Q4

4.1 1

(\bar{X}) برای فرایند داده شده داریم:

$$\begin{aligned}\mathbb{E}[X(t)] &= \mathbb{E}[X_1(t)] + \mathbb{E}[cX_2(t)] \\ &= \mathbb{E}[X_1(t)] + \mathbb{E}[c] \mathbb{E}[X_2(t)] \\ &= \eta_1 + 0.5\eta_2\end{aligned}$$

این در حالی است که برای sample path ای که $c=0$ می باشد $X(t) = X_1(t)$ می باشد که در نتیجه وقتی $T \rightarrow \infty$ میانگین بدست آمده برابر $\eta_1 \rightarrow \eta_T$ خواهد بود. بنابراین فرایند داده شده Mean Ergodic نمی باشد.

4.2 2

$$\begin{aligned}\eta_t &= \frac{1}{2T} \int_{-T}^T (a \cos(\omega_1 t) + b \cos(\omega_2 t) + c) dt = \\ &= \frac{a}{2T\omega_1} \sin(\omega_1 t) \Big|_{t=-T}^T + \frac{b}{2T\omega_2} \sin(\omega_2 t) \Big|_{t=-T}^T + \frac{c}{2T} t \Big|_{t=-T}^T \\ &= \frac{a}{\omega_1 T} \sin(\omega_1 T) + \frac{b}{\omega_2 T} \sin(\omega_2 T) + c \\ \text{Var}(\eta_t) &\Rightarrow \mathbb{E} \left[\left(\frac{a}{\omega_1 T} \sin(\omega_1 T) + \frac{b}{\omega_2 T} \sin(\omega_2 T) + c - c \right)^2 \right] \\ T \rightarrow \infty & \\ \lim_{T \rightarrow \infty} &= \mathbb{E} \left[\frac{a^2}{\omega_1^2 T^2} \sin^2(\omega_1 T) + \frac{b^2}{\omega_2^2 T^2} \sin^2(\omega_2 T) + \frac{2ab}{\omega_1 \omega_2 T^2} \sin(\omega_1 T) \sin(\omega_2 T) \right] = 0\end{aligned}$$

Mean Ergodic

4.3 3

(ج) مشابه بخش قبل ابتدا WSS بودن یا نبودن فرایند را بررسی می‌کنیم:

$$\mathbb{E}[X(t)] = A\mathbb{E}[\cos(\omega t + \phi)] = \frac{1}{2\pi} \int_0^{2\pi} \cos(\omega t + \theta) d\theta = 0$$

$$\begin{aligned} \mathbb{E}[X(t_1)X(t_2)] &= A^2\mathbb{E}[\cos(\omega t_1 + \phi)\cos(\omega t_2 + \phi)] \\ &= \frac{A^2}{2}\mathbb{E}[\cos(\omega(t_1 - t_2)) + \cos(\omega(t_1 + t_2) + 2\phi)] \\ &= \frac{A^2}{2}[\mathbb{E}[\cos(\omega(t_1 - t_2))] + \mathbb{E}[\cos(\omega(t_1 + t_2) + 2\phi)]] \\ &= \frac{A^2}{2}\cos(\omega(t_1 - t_2)) \end{aligned}$$

بنابراین فرایند WSS می‌باشد. حال طبق قضیه Slutsky داریم:

$$\begin{aligned} \lim_{T \rightarrow \infty} \frac{1}{T} \int_0^T C(\tau) d\tau &= \lim_{T \rightarrow \infty} \frac{1}{T} \int_0^T \frac{A^2}{2} \cos(\omega\tau) d\tau \\ &= \lim_{T \rightarrow \infty} \frac{A^2}{2\omega T} \sin(\omega T) \\ &= 0 \end{aligned}$$

بنابراین فرایند Mean Ergodic می‌باشد.

5 Q5

$$\underline{y}(t) = 2\underline{x}(t) + 3\underline{x}'(t) \quad \eta_x = 5 \quad C_{xx}(\tau) = 4e^{-2|\tau|}$$

The process $\underline{y}(t)$ is the output of the system $H(s) = 2+3s$ with input $\underline{x}(t)$. Hence,
 $\eta_y = 5H(0) = 10$

$$S_{yy}^c(\omega) = S_{xx}^c(\omega)|2+3j\omega|^2 = \frac{16}{4+\omega^2}(4+9\omega^2) = 144 - \frac{512}{4+\omega^2} = S_{yy}(\omega) - 2\pi\eta_y^2\delta(\omega)$$

$$S_{yy}(\omega) = 144 - \frac{512}{4+\omega^2} + 2\pi(10)^2\delta(\omega)$$

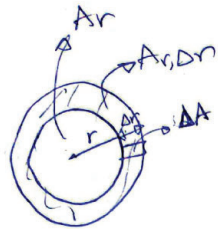
$$R_{yy}(\tau) = 144\delta(\tau) - \frac{128}{2^2+\omega^2} \frac{2(\tau)}{2^2+\omega^2}$$

$$S_{yy}(\omega) = 144 - 128 \frac{2(\tau)}{2^2+\omega^2} + 2\pi(100)\delta(\omega)$$

$$= 144\delta(\tau) - 128 e^{-a|\tau|} + 100$$

$$= 288\pi\delta(\tau) - 128 e^{-a|\tau|} + 100$$

6 Q6



پیش ششم :

(۱) نزدیک ترین نقطه

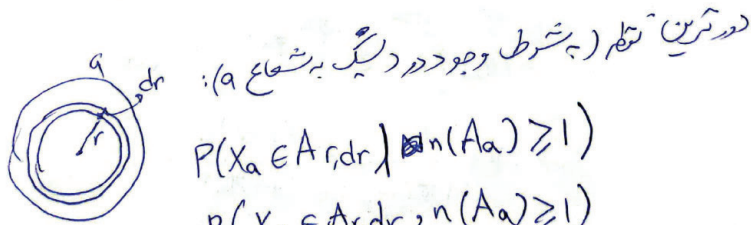
$n(A)$ - تعداد نقاط داخل A

$$P(X_0 \in A_{r, \Delta r}) = P(n(A_r) = 0, n(A_{r, \Delta r}) \geq 1) \\ = e^{-\lambda S(A_r)} (1 - e^{-\lambda S(A_{r, \Delta r})})$$

$$\rightarrow P(X_0 \in \Delta A) = \frac{\Delta A}{S(A_{r, \Delta r})} e^{-\lambda S(A_r)} (1 - e^{-\lambda S(A_{r, \Delta r})})$$

$$\rightarrow \lim_{\Delta A \rightarrow 0} \frac{P(X_0 \in \Delta A)}{\Delta A} = \lim_{\Delta A \rightarrow 0} \frac{\lambda e^{-\lambda \pi r^2} (1 - e^{-\lambda S(A_{r, \Delta r})})}{S(A_{r, \Delta r})}$$

$$\Rightarrow \lambda e^{-\lambda \pi r^2} \rightarrow f(x_0) = e^{-\lambda \pi r(x_0)^2}$$



دورترین نقطه (به شرط وجود در یک شعاع a):

$$P(X_a \in A_{r, dr} | n(A_a) \geq 1)$$

$$= \frac{P(X_a \in A_{r, dr}, n(A_a) \geq 1)}{P(n(A_a) \geq 1)}$$

$$= \frac{(1 - e^{-\lambda S(A_{r, dr})}) \cdot e^{-\lambda S(A_{r, dr}, a)}}{1 - e^{-\lambda S(A_a)}}$$

Scanned with CamScanner

$$\rightarrow f(x_a | n(A_a) \geq 1) = \frac{\lambda e^{-\lambda \pi (a^2 - r^2)}}{1 - e^{-\lambda \pi a^2}}$$

~~$Cor(X, X_0) = E[X X_0] - E[X] E[X_0]$~~

$$Cor(N(x), N(x_0)) = E[N(x)N(x_0)]$$

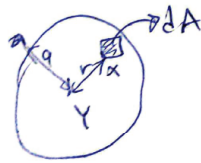
$$= E[N(x)] E[N(x_0)]$$

$$= \frac{r\pi\lambda e^{-\lambda\pi a^2}}{1 - e^{-\lambda\pi a^2}} \int_0^a e^{-\left(\frac{1}{4r} - \lambda\pi\right)r^2} 2\pi r dr$$

$$= \frac{r\pi\lambda e^{-\lambda\pi a^2}}{1 - e^{-\lambda\pi a^2}} \cdot \frac{1}{\frac{1}{4r} - \lambda\pi} \left(1 - e^{-\left(\frac{1}{4r} - \lambda\pi\right)a^2}\right)$$

$$= \frac{r\pi\lambda}{\left(\frac{1}{4r} - \lambda\pi\right)(1 - e^{-\lambda\pi a^2})} \left(e^{-\lambda\pi a^2} - e^{-\frac{a^2}{4r}}\right)$$

یا ابتدا دست کنید که \tilde{w} ب M توزیع پواسن با $\lambda = \frac{1}{\pi}$ است



$$E[m(Y)]$$

$$= \lim_{dA \rightarrow 0} \sum_{dA} P(Y \in B_x(r(x))) \lambda \pi (dA)$$

$$\times E[N(Y)N(x)]$$

$$= \lim_{dA \rightarrow 0} \sum_{dA} \left(1 - \frac{r}{a}\right) \frac{\lambda}{r} dA \times \frac{1}{1+r}$$

$$\Rightarrow = \int_{B_0(Y)} \left(1 - \frac{r}{a}\right) \frac{\lambda}{r} \cdot \frac{1}{1+r} dA$$

$$= \int_0^{2\pi} \int_0^a \left(1 - \frac{r}{a}\right) \frac{\lambda}{r} \cdot \frac{1}{1+r} r dr d\theta$$

$$\begin{aligned}
&= \frac{r\pi\lambda}{r_a} \int_1^{a+1} -\frac{a+1}{r} \cdot -r + (a+r) \, dr \\
&= \frac{r\pi\lambda}{r_a} \left(-(a+1)\ln(a+1) - \frac{(a+1)^r - 1}{r} + (a+r)a \right) \\
&= \frac{r\pi\lambda}{r} \left(-\left(1 + \frac{1}{a}\right)\ln(a+1) - \frac{(a+r)}{r} \right)
\end{aligned}$$